

Homework assignment 2 – Applied Statistics 2018

Hand in **your own solutions** at the start of the lecture on 23/3 or (Kanifing, 11.00 h)).

A random variable X has a density function f with unknown parameter $\theta > 0$, defined as follows:

$$f(x) = \begin{cases} \frac{2x}{\theta^2}, & \text{if voor } 0 \leq x \leq \theta \\ 0, & \text{elsewhere} \end{cases}$$

A random sample X_1, \dots, X_n of X is available.

- Determine the first three moments $E(X)$, $E(X^2)$, $E(X^3)$ and $\text{var}(X)$
- Sketch the graph of the density function, state whether the distribution is skewed (to the left or the right?) and verify that the skewness coefficient of the distribution has a corresponding value (use a.)
- Use the result of a. to determine the moment estimator of θ .
- Show that $\frac{3}{2}\bar{X}$ is an unbiased and consistent estimator of θ .
- Consider all estimators $T = a \cdot \bar{X}$, with real constant $a \in \mathbb{R}$.
For which value of a is T the best estimator of θ .
- Determine the maximum likelihood estimator of θ .

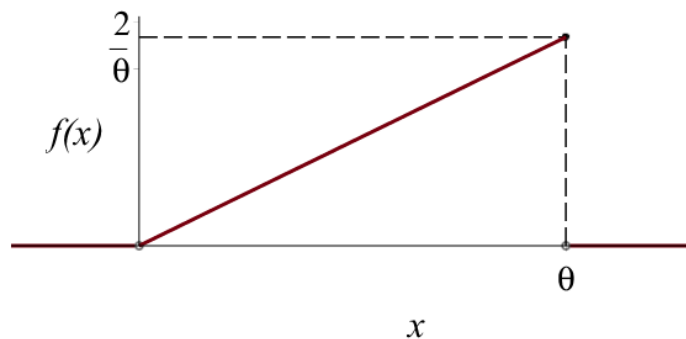
a	b	c	d	e	f	Total
2	1.5	1	2	2	1.5	10

Solutions:

$$\begin{aligned} \text{a. } E(X) &= \int_{-\infty}^{\infty} xf(x)dx = \int_0^{\theta} x \cdot \frac{2x}{\theta^2} dx = \left[\frac{2x^3}{3\theta^2} \right]_{x=0}^{\theta} = \frac{2}{3}\theta \\ E(X^2) &= \int_{-\infty}^{\infty} x^2 f(x)dx = \int_0^{\theta} x^2 \cdot \frac{2x}{\theta^2} dx = \left[\frac{2x^4}{4\theta^2} \right]_{x=0}^{\theta} = \frac{1}{2}\theta^2 \\ E(X^3) &= \int_{-\infty}^{\infty} x^3 f(x)dx = \int_0^{\theta} x^3 \cdot \frac{2x}{\theta^2} dx = \left[\frac{2x^5}{5\theta^2} \right]_{x=0}^{\theta} = \frac{2}{5}\theta^3 \\ \text{var}(X) &= E(X^2) - (EX)^2 = \frac{1}{2}\theta^2 - \frac{4}{9}\theta^2 = \frac{1}{18}\theta^2 \end{aligned}$$

- b.** The graph is skewed to the left (“tail” on the left hand side).

Graph of the density function f



The skewness coefficient is $\gamma_1 = E(X - \mu)^3 / \sigma^3$.

In a. we found $\sigma^2 = \frac{1}{18}\theta^2$ ($\sigma = \frac{\theta}{\sqrt{18}}$).

$$\begin{aligned} E(X - \mu)^3 &= E[X^3 - 3\mu X^2 + 3\mu^2 X - \mu^3] = E(X^3) - 3\mu \cdot E(X^2) + 3\mu^2 \cdot E(X) - \mu^3 \\ &= E(X^3) - 3\mu \cdot E(X^2) + 2\mu^3 \stackrel{\text{a.}}{=} \frac{2}{5}\theta^3 - 3 \cdot \frac{1}{3}\theta^3 + 2 \cdot \frac{8}{27}\theta^3 = -\frac{1}{135}\theta^3 \end{aligned}$$

$$\gamma_1 = \frac{E(X-\mu)^3}{\sigma^3} = \frac{-\frac{1}{135}\theta^3}{(\theta/\sqrt{18})^3} = -\frac{2}{5}\sqrt{2} \approx -0.56 \quad (< 0, \text{ as expected}).$$

- c.** $E(X) = \frac{2}{3}\theta$, so $\theta = \frac{3}{2}\mu$ and \bar{X} is the moment estimator of μ : $\frac{3}{2}\bar{X}$ is the moment estimator of θ .

d. $E(\bar{X}) = \mu = \frac{2}{3}\theta$ and $var(\bar{X}) = \frac{\sigma^2}{n} = \frac{\theta^2}{18n}$

$E\left(\frac{3}{2}\bar{X}\right) = \frac{3}{2}E(\bar{X}) = \frac{3}{2} \cdot \frac{2}{3}\theta = \theta \Rightarrow \frac{3}{2}\bar{X}$ is an unbiased estimator of θ .

Furthermore $var\left(\frac{3}{2}\bar{X}\right) = \frac{9}{4}var(\bar{X}) = \frac{\theta^2}{8n}$ and $\lim_{n \rightarrow \infty} var\left(\frac{3}{2}\bar{X}\right) = \lim_{n \rightarrow \infty} \frac{\theta^2}{8n} = 0$,

so unbiased and variance converges to 0, then $\frac{3}{2}\bar{X}$ is a consistent estimator of θ .

e. $E(T) = a \cdot E(\bar{X}) = \frac{2a\theta}{3}$ and $var(T) = a^2 \cdot var(\bar{X}) = \frac{a^2\theta^2}{18n}$

$MSE(T) = (ET - \theta)^2 + var(T) = \left(\frac{2a}{3} - 1\right)^2 \theta^2 + \frac{a^2\theta^2}{18n} = \theta^2 \left[\left(\frac{2a}{3} - 1\right)^2 + \frac{a^2}{18n} \right]$

Determine maximum of $f(a) = \left(\frac{2a}{3} - 1\right)^2 + \frac{a^2}{18n}$:

$f'(a) = \frac{4}{3}\left(\frac{2a}{3} - 1\right) + \frac{a}{9n} = 0 \Leftrightarrow \left(\frac{8}{9} + \frac{1}{9n}\right)a - \frac{4}{3} = 0 \Leftrightarrow a = \frac{4}{3} \cdot \frac{8n}{8n+1}$

Since $f''(a) = \frac{8}{9} + \frac{1}{n} > 0$, $f(a)$ attains a minimum at $a = \frac{12n}{8n+1}$

$T = \frac{12n}{8n+1}\bar{X}$ is the best estimator (note that $E(T) = \frac{12n}{8n+1} \cdot \frac{2}{3}\theta$ is not unbiased, but

$\lim_{n \rightarrow \infty} E(T) = \lim_{n \rightarrow \infty} \frac{12n}{8n+1} \cdot \frac{2}{3}\theta = \frac{12}{8} \cdot \frac{2}{3}\theta = \theta$:

$T = \frac{12n}{8n+1}\bar{X}$ is asymptotically unbiased).

f. $L(\theta) = \prod f(x_i|\theta) = 2^n(x_1 \cdot \dots \cdot x_n) \cdot \frac{1}{\theta^n}$, where $0 \leq x_i \leq \theta$ for all $i = 1, \dots, n$,

so domain of L is: $\theta \geq \max(x_1, \dots, x_n)$.

L is a decreasing function in θ ($L'(\theta) = -n2^n(x_1 \cdot \dots \cdot x_n) \cdot \frac{1}{\theta^{n+1}} < 0$) and attains its maximum in the bound $\theta = \max(x_1, \dots, x_n)$.

Hence $\hat{\theta} = \max(X_1, \dots, X_n)$ is the maximum likelihood estimator of θ .